

# Identifying or Verifying the Number of Factors to Extract using Very Simple Structure.

As published in Benchmarks RSS Matters, December 2014

<http://web3.unt.edu/benchmarks/issues/2014/12/rss-matters>

Jon Starkweather, PhD

Jon Starkweather, PhD  
jonathan.starkweather@unt.edu  
Consultant  
**Research and Statistical Support**



<http://www.unt.edu>



<http://www.unt.edu/rss>

RSS hosts a number of “Short Courses”.  
A list of them is available at:  
<http://www.unt.edu/rss/Instructional.htm>

Those interested in learning more about R, or how to use it, can find information here:  
[http://www.unt.edu/rss/class/Jon/R\\_SC](http://www.unt.edu/rss/class/Jon/R_SC)

# Identifying or Verifying the Number of Factors to Extract using Very Simple Structure.

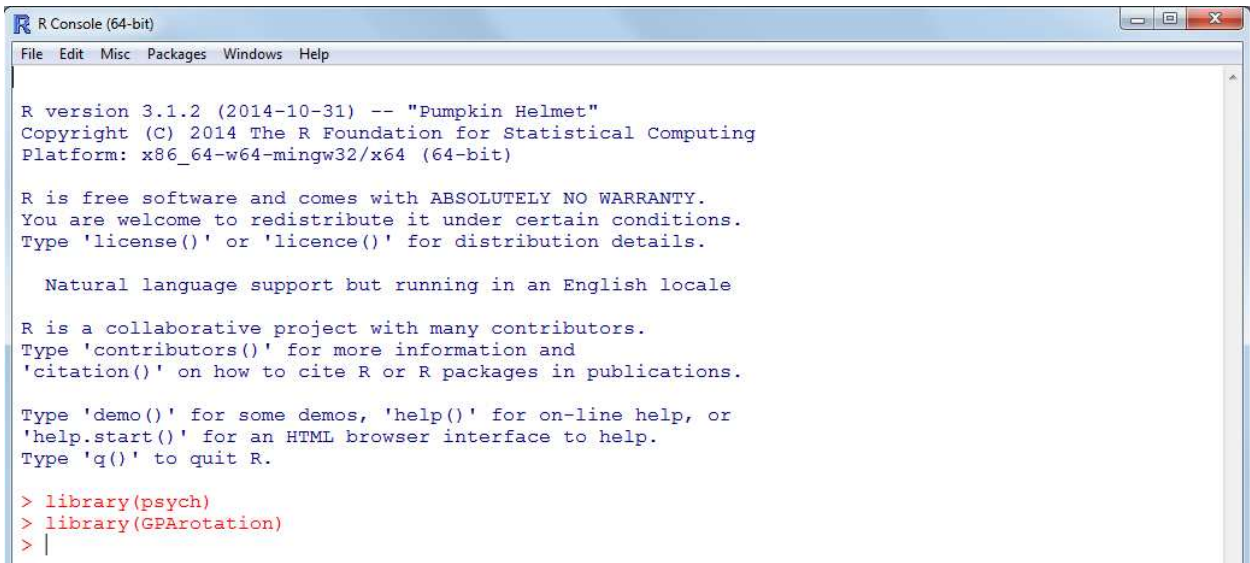
Factor analysis is perhaps one of the most frequently used analyses. It is versatile and flexible; meaning, it can be applied to a variety of data situations and types, and it can be applied in a variety of ways. However, conducting factor analysis generally requires the data analyst to make several decisions. Analysts often run several factor analyses, even when attempting to confirm an established factor structure; in order to assess the fit of the data to several factor models (e.g. one factor model, two factor model, three factor model, etc.). Over the 100 years since Spearman (1904) developed factor analysis there have been many, many criteria proposed for determining the number of factors to extract (e.g. eigenvalues greater than one, Horn's [1965] parallel analysis, Cattell's [1966] scree plot or test, Velicer's [1976] Minimum Average Partial [MAP] criterion, etc.). Each of these proposed criteria have strengths and weaknesses; and they occasionally conflict with one another, which makes using one criterion over another a risky proposition. This month's article demonstrates a very handy method for comparing multiple criteria in the pursuit of choosing to extract the appropriate number of factors during factor analysis.

In popular culture it is not uncommon to hear someone say, "There's an *app* for that." The phrase generally refers to the idea that an *application* exists (for a smart phone) which does the task being discussed. Likewise, here at RSS we very frequently find "There's a *pack* for that." This phrase refers to the virtual certainty of finding an R *package* which has a function devoted to some analysis or technique we are discussing. The primary package we will be using here is one package which contains a great many useful functions and as a result is very often *the* package we end up using for a variety of analyses. The primary package we will be using here is the 'psych' package (Revelle, 2014). The 'psych' package has grown substantially over the last few years and includes many very useful functions — if you have not taken a look at it recently, you might want to check it out.

Our examples below will actually require two packages, the 'psych' package and the 'GPArotation' package (Bernaards & Jennrich, 2014). The 'GPArotation' package should be familiar to anyone with experience doing factor analysis — it provides functions for several rotation strategies. The primary function we demonstrate below is the 'vss' function from the 'psych' package. The *Very Simple Structure* (VSS; Revell & Rocklin, 1979) function provides a nice output of criteria for varying levels of factor model complexity (i.e. number of factors to extract). The Very Simple Structure (VSS) terminology is used to refer to the idea that all loadings which are less than the maximum loading (of an item to a factor) are suppressed to zero — thus forcing a particular factor model to be much more interpretable or more clearly distinguished. Then, fit of several models of increasing rank complexity (i.e. more and more factors specified) can be assessed using the residual matrix of each model (i.e. original matrix minus the reproduced matrix of the models). We will also be using both the 'fa' function (from the 'psych' package) and the 'factanal' function (from the 'stats' package — included with all installations of R) to fit factor analysis models to the data structures.

## 1 Examples

The first two examples used here can easily be duplicated using the scripts provided below (i.e. the data file is available at the URL in the script / screen capture image). The third example is the example contained in the help file of the 'vss' function and can be accessed using the script below. First, load the two packages we will be using.



```
R Console (64-bit)
File Edit Misc Packages Windows Help

R version 3.1.2 (2014-10-31) -- "Pumpkin Helmet"
Copyright (C) 2014 The R Foundation for Statistical Computing
Platform: x86_64-w64-mingw32/x64 (64-bit)

R is free software and comes with ABSOLUTELY NO WARRANTY.
You are welcome to redistribute it under certain conditions.
Type 'license()' or 'licence()' for distribution details.

  Natural language support but running in an English locale

R is a collaborative project with many contributors.
Type 'contributors()' for more information and
'citation()' on how to cite R or R packages in publications.

Type 'demo()' for some demos, 'help()' for on-line help, or
'help.start()' for an HTML browser interface to help.
Type 'q()' to quit R.

> library(psych)
> library(GPArotation)
> |
```

Next, we will import the comma delimited text (.txt) file from the RSS server using the URL and file name (vss\_df.txt) contained in the script / image below. We also run a simple 'summary' on the data frame to make sure it was imported correctly.

```

R Console (64-bit)
File Edit Misc Packages Windows Help

> vss.df <- read.table("http://www.unt.edu/rss/class/Jon/Benchmarks/vss_df.txt",
+   header = TRUE, sep = ",", na.strings = "NA", dec = ".", strip.white = TRUE)
> summary(vss.df)

```

s.id	group	age	sex	class	i1
Min. : 1	Min. :1.000	Min. :18.00	Female:908	Freshman :386	Min. : -3.095269
1st Qu.:2352	1st Qu.:1.000	1st Qu.:21.00	Male :492	Junior :371	1st Qu.: -0.610912
Median :4630	Median :2.000	Median :24.00		Senior :254	Median : -0.004898
Mean :4578	Mean :1.701	Mean :24.23		Sophomore:389	Mean : 0.010447
3rd Qu.:6871	3rd Qu.:2.000	3rd Qu.:27.00			3rd Qu.: 0.703460
Max. :8999	Max. :2.000	Max. :32.00			Max. : 3.558348

i2	i3	i4	i5	i6
Min. : -2.774878	Min. : -4.249428	Min. : -3.615968	Min. : -3.19226	Min. : -3.06995
1st Qu.: -0.687474	1st Qu.: -0.668724	1st Qu.: -0.683154	1st Qu.: -0.68723	1st Qu.: -0.63830
Median : 0.007035	Median : -0.052491	Median : 0.031426	Median : 0.02228	Median : -0.01050
Mean : 0.022828	Mean : -0.000987	Mean : 0.009961	Mean : 0.01117	Mean : 0.02408
3rd Qu.: 0.098613	3rd Qu.: 0.682047	3rd Qu.: 0.680406	3rd Qu.: 0.71961	3rd Qu.: 0.68564
Max. : 3.081032	Max. : 3.436386	Max. : 3.175091	Max. : 3.40557	Max. : 3.16402

i7	i8	i9	i10	i11
Min. : -3.343048	Min. : -3.37473	Min. : -3.44980	Min. : -3.31567	Min. : -3.31370
1st Qu.: -0.666585	1st Qu.: -0.72200	1st Qu.: -0.65198	1st Qu.: -0.66184	1st Qu.: -0.63443
Median : -0.014271	Median : -0.03628	Median : 0.03575	Median : 0.01875	Median : 0.04452
Mean : 0.008293	Mean : -0.02285	Mean : 0.02056	Mean : 0.00282	Mean : 0.02508
3rd Qu.: 0.632732	3rd Qu.: 0.65504	3rd Qu.: 0.68236	3rd Qu.: 0.67405	3rd Qu.: 0.69187
Max. : 3.599682	Max. : 3.24299	Max. : 3.35574	Max. : 3.24593	Max. : 3.35810

i12	i13	i14	i15	i16
Min. : -3.86434	Min. : -4.133794	Min. : -3.29740	Min. : -3.097888	Min. : -3.74943
1st Qu.: -0.71402	1st Qu.: -0.699098	1st Qu.: -0.71330	1st Qu.: -0.683492	1st Qu.: -0.68303
Median : -0.02162	Median : -0.002602	Median : 0.01323	Median : -0.002551	Median : -0.03982
Mean : -0.02977	Mean : -0.012313	Mean : -0.01453	Mean : -0.006453	Mean : -0.02930
3rd Qu.: 0.65953	3rd Qu.: 0.664155	3rd Qu.: 0.67924	3rd Qu.: 0.660344	3rd Qu.: 0.61757
Max. : 2.93930	Max. : 3.090504	Max. : 3.08462	Max. : 3.370616	Max. : 2.82001

i17	i18	i19	i20	i21
Min. : -3.047998	Min. : -2.77885	Min. : -3.05053	Min. : -3.14663	Min. : -3.0113
1st Qu.: -0.718641	1st Qu.: -0.70987	1st Qu.: -0.72716	1st Qu.: -0.70180	1st Qu.: -0.7058
Median : 0.011685	Median : -0.02235	Median : -0.04124	Median : -0.06383	Median : -0.0512
Mean : -0.006581	Mean : -0.03562	Mean : -0.01607	Mean : -0.02129	Mean : -0.0253
3rd Qu.: 0.667417	3rd Qu.: 0.62268	3rd Qu.: 0.67362	3rd Qu.: 0.63290	3rd Qu.: 0.6451
Max. : 3.184236	Max. : 2.97589	Max. : 3.28403	Max. : 3.45381	Max. : 3.1598

i22	i23	i24	i25	i26
Min. : -3.57302	Min. : -3.43963	Min. : -3.09786	Min. : -3.84431	Min. : -3.483726
1st Qu.: -0.63999	1st Qu.: -0.73039	1st Qu.: -0.68945	1st Qu.: -0.72019	1st Qu.: -0.711454
Median : -0.01984	Median : -0.05164	Median : -0.02866	Median : -0.06525	Median : -0.003657
Mean : -0.02366	Mean : -0.03144	Mean : -0.01272	Mean : -0.04618	Mean : 0.003004
3rd Qu.: 0.60841	3rd Qu.: 0.66109	3rd Qu.: 0.64324	3rd Qu.: 0.61634	3rd Qu.: 0.699988
Max. : 3.30948	Max. : 4.03608	Max. : 3.99969	Max. : 3.11480	Max. : 2.927518

i27	i28	i29	i30
Min. : -3.82320	Min. : -3.316482	Min. : -3.631834	Min. : -3.70954
1st Qu.: -0.67767	1st Qu.: -0.665499	1st Qu.: -0.696295	1st Qu.: -0.69770
Median : 0.00007	Median : -0.028020	Median : 0.038772	Median : -0.04353
Mean : 0.01349	Mean : 0.000607	Mean : 0.005513	Mean : -0.02384
3rd Qu.: 0.66166	3rd Qu.: 0.661875	3rd Qu.: 0.698633	3rd Qu.: 0.66926
Max. : 3.10125	Max. : 3.175330	Max. : 3.358999	Max. : 3.14729

```

> |

```

The simulated data includes a sample identification number for each participant (s.id), a grouping variable (group 1 or group 2), age of each participant (age in years), sex of each participant (female or male), class standing of each participant (freshman, sophomore, junior, or senior), and 30 item scores. Next, we will identify which participants belong to group 1 and which belong to group 2; as well as the number of participants in each group.

```

R Console (64-bit)
File Edit Misc Packages Windows Help

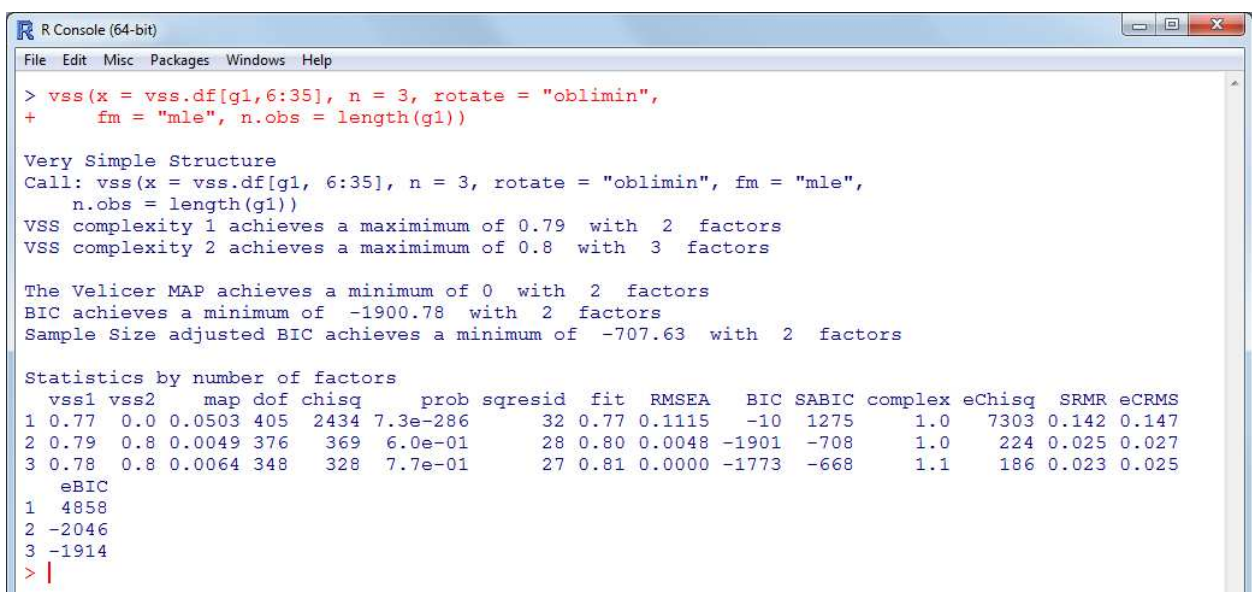
> g1 <- which(vss.df[,2] == 1); length(g1)
[1] 418
> g2 <- which(vss.df[,2] == 2); length(g2)
[1] 982
> |

```

So, we have 418 participants in group 1 and 982 participants in group 2. Generally when analysts intend



to do factor analysis they have an idea of how many factors they believe the appropriate factor model contains; and often they have an idea of whether an orthogonal or oblique rotation strategy is warranted. For this first example (i.e. group 1) looking at the 30 item scores (i.e. columns 6 through 35), we believe there are two factors and therefore; we specify 3 factors ( $n = 3$ ) in the ‘vss’ function. We also believe the factors are likely to be meaningfully related and consequently, we specify an oblimin rotation strategy. Next, we apply the ‘vss’ function to group 1. Also note, we specified Maximum Likelihood Estimation as the Factor Method ( $fm = "mle"$ ) because this is the method used by default with the ‘factanal’ (i.e. factor analysis) function of the ‘stats’ package. We specified the number of observations (i.e. number of rows, cases, or participants) using the length of the group 1 vector ( $g1$ ). Recall from above, the group 1 vector contains the row numbers of all the participants from group 1.



```

R Console (64-bit)
File Edit Misc Packages Windows Help

> vss(x = vss.df[g1,6:35], n = 3, rotate = "oblimin",
+     fm = "mle", n.obs = length(g1))

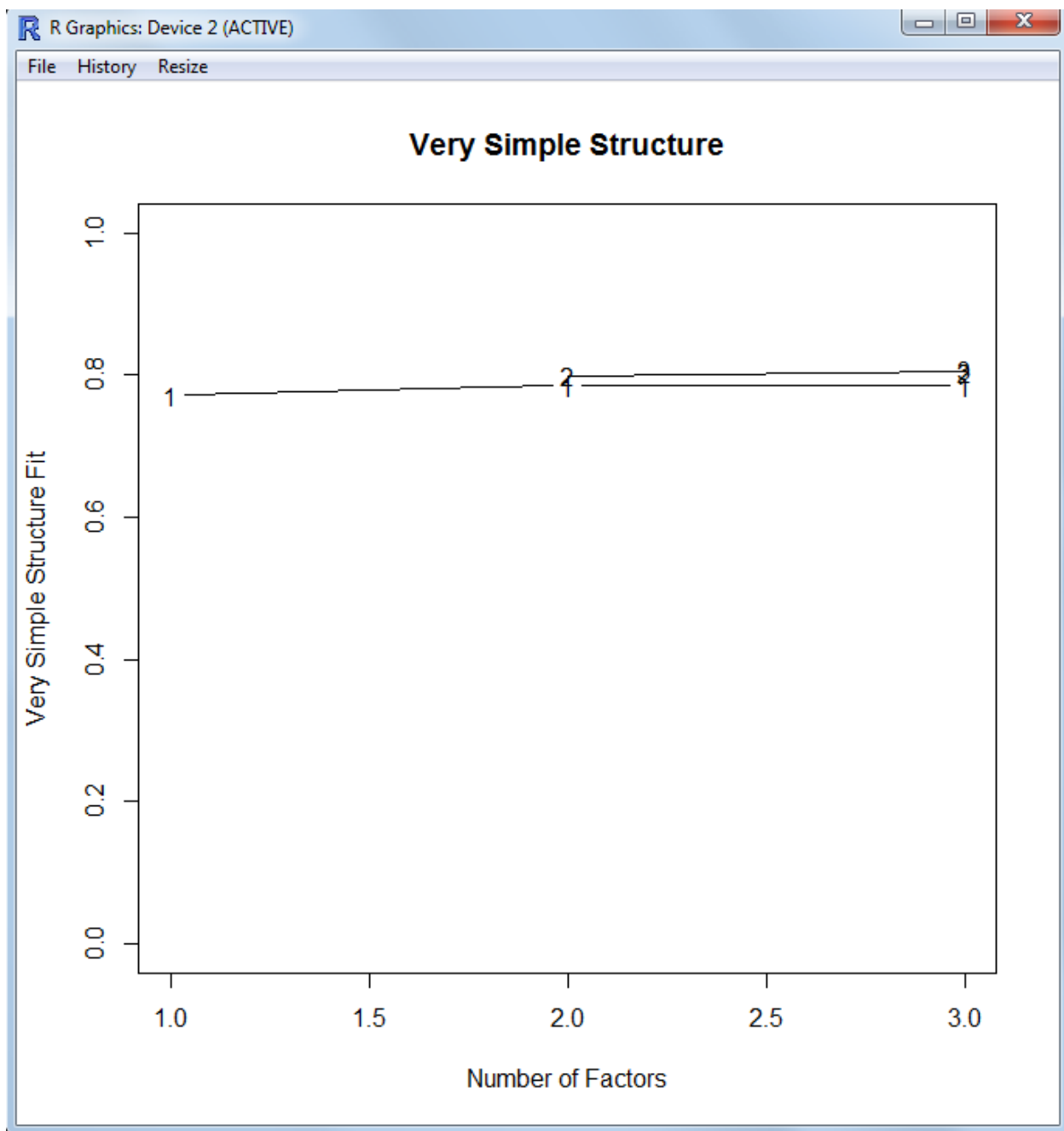
Very Simple Structure
Call: vss(x = vss.df[g1, 6:35], n = 3, rotate = "oblimin", fm = "mle",
  n.obs = length(g1))
VSS complexity 1 achieves a maximum of 0.79 with 2 factors
VSS complexity 2 achieves a maximum of 0.8 with 3 factors

The Velicer MAP achieves a minimum of 0 with 2 factors
BIC achieves a minimum of -1900.78 with 2 factors
Sample Size adjusted BIC achieves a minimum of -707.63 with 2 factors

Statistics by number of factors
  vss1 vss2  map dof chisq  prob sqresid  fit  RMSEA  BIC  SABIC complex eChisq  SRMR eCRMS
1 0.77 0.0 0.0503 405 2434 7.3e-286 32 0.77 0.1115 -10 1275 1.0 7303 0.142 0.147
2 0.79 0.8 0.0049 376 369 6.0e-01 28 0.80 0.0048 -1901 -708 1.0 224 0.025 0.027
3 0.78 0.8 0.0064 348 328 7.7e-01 27 0.81 0.0000 -1773 -668 1.1 186 0.023 0.025
  eBIC
1 4858
2 -2046
3 -1914
>

```

The first few rows of output (i.e. “Very Simple Structure” table) show the function called and the *maximum* complexity values. This is a good example because the VSS complexity rows are conflicting; VSS complexity 1 shows a 2–factor model is best while VSS complexity 2 indicates a 3–factor model is best. The VSS complexity 2 is a bit misleading because both the 2–factor model and 3–factor model display a VSS complexity 2 of 0.80; as can be seen in the first column of output under the “Statistics by number of factors” table. So, in fact both complexity 1 and complexity 2 are in agreement. Furthermore, the Velicer MAP *minimum* is reached with the 2–factor model; which can also be seen in the third column of the “Statistics by number of factors” table. The Bayesian Information Criterion (BIC) *minimum* is reached with the 2–factor model; as well as the Sample Size adjusted BIC (SABIC) — shown in columns 10 and 11 respectively of the “Statistics by number of factors” table. The ‘vss’ function also produces a plot (by default) which shows the number of factors on the x–axis and the VSS (complexity) Fit along the y–axis with lines and numbers in the Cartesian plane representing the (3) different factor models (see below).



To interpret the graph, focus on the model (1, 2, or 3 factor models) which has the highest line (and numerals) in relation to the y-axis; but also note any transitions of the model lines. In this example, the transitions are all very nearly flat but a later example will better demonstrate the utility of this type of plot.

Next, we can verify the fit of our 2-factor model using either the 'fa' function (from the 'psych' package) and / or the 'factanal' function (of the 'stats' package).

```

R Console (64-bit)
File Edit Misc Packages Windows Help

> fa(r = vss.df[g1,6:35], nfactors = 2, rotate = "oblimin", fm = "mle")
Factor Analysis using method = ml
Call: fa(r = vss.df[g1, 6:35], nfactors = 2, rotate = "oblimin", fm = "mle")
Standardized loadings (pattern matrix) based upon correlation matrix
      ML1    ML2    h2    u2 com
i1  0.89  0.03  0.82  0.18  1.0
i2  0.82 -0.04  0.65  0.35  1.0
i3  0.83 -0.01  0.68  0.32  1.0
i4  0.49 -0.01  0.23  0.77  1.0
i5  0.72 -0.01  0.51  0.49  1.0
i6  0.65  0.01  0.43  0.57  1.0
i7  0.65  0.00  0.43  0.57  1.0
i8  0.48  0.12  0.29  0.71  1.1
i9  0.65  0.00  0.42  0.58  1.0
i10 0.66  0.03  0.45  0.55  1.0
i11 0.80  0.03  0.66  0.34  1.0
i12 0.53 -0.01  0.28  0.72  1.0
i13 0.62  0.01  0.39  0.61  1.0
i14 0.69 -0.03  0.46  0.54  1.0
i15 0.84 -0.02  0.69  0.31  1.0
i16 -0.01  0.90  0.80  0.20  1.0
i17 0.04  0.74  0.58  0.42  1.0
i18 -0.01  0.79  0.62  0.38  1.0
i19 -0.03  0.52  0.26  0.74  1.0
i20 0.02  0.65  0.44  0.56  1.0
i21 -0.03  0.58  0.33  0.67  1.0
i22 0.02  0.55  0.31  0.69  1.0
i23 0.01  0.39  0.16  0.84  1.0
i24 0.00  0.67  0.45  0.55  1.0
i25 -0.06  0.70  0.46  0.54  1.0
i26 0.04  0.76  0.61  0.39  1.0
i27 0.11  0.38  0.20  0.80  1.2
i28 0.07  0.54  0.33  0.67  1.0
i29 -0.04  0.68  0.44  0.56  1.0
i30 0.01  0.80  0.65  0.35  1.0

      ML1    ML2
SS loadings      7.39 6.60
Proportion Var    0.25 0.22
Cumulative Var    0.25 0.47
Proportion Explained 0.53 0.47
Cumulative Proportion 0.53 1.00

With factor correlations of
      ML1    ML2
ML1 1.00 0.44
ML2 0.44 1.00

Mean item complexity = 1
Test of the hypothesis that 2 factors are sufficient.

The degrees of freedom for the null model are 435 and the objective function was 16.06 with Chi$
The degrees of freedom for the model are 376 and the objective function was 0.91

```

Note: the last few lines of output from the 'fa' function are cut off (i.e. not shown).



```

R Console (64-bit)
File Edit Misc Packages Windows Help

> factanal(vss.df[g1,6:35], factors = 2, rotation = "oblimin")

Call:
factanal(x = vss.df[g1, 6:35], factors = 2, rotation = "oblimin")

Uniquenesses:
  i1  i2  i3  i4  i5  i6  i7  i8  i9  i10 i11 i12 i13 i14 i15 i16
0.178 0.351 0.325 0.768 0.492 0.574 0.573 0.706 0.580 0.547 0.337 0.723 0.608 0.542 0.313 0.200
  i17 i18 i19 i20 i21 i22 i23 i24 i25 i26 i27 i28 i29 i30
0.422 0.379 0.742 0.565 0.675 0.694 0.841 0.554 0.545 0.390 0.804 0.665 0.558 0.355

Loadings:
  Factor1 Factor2
i1  0.891
i2  0.823
i3  0.828
i4  0.487
i5  0.717
i6  0.647
i7  0.653
i8  0.478  0.123
i9  0.647
i10 0.660
i11 0.801
i12 0.530
i13 0.621
i14 0.689
i15 0.838
i16      0.899
i17      0.744
i18      0.793
i19      0.521
i20      0.649
i21      0.584
i22      0.547
i23      0.394
i24      0.669
i25      0.700
i26      0.762
i27 0.110  0.384
i28      0.542
i29      0.680
i30      0.799

SS loadings      Factor1 Factor2
Proportion Var   0.245  0.219
Cumulative Var   0.245  0.464

Factor Correlations:
  Factor1 Factor2
Factor1  1.000 -0.435
Factor2 -0.435  1.000

```

Note: last few lines of output from the ‘factanal’ function are cut off (i.e. not shown).

We will now assess the group 2 (g2) data. This group is believed to be best served with a 3–factor model; so we specify 4 factors ( $n = 4$ ) in the ‘vss’ function call; again with the factor method set to Maximum Likelihood Estimation (fm = “mle”) and an oblique rotation strategy (rotate = “oblimin”).

```

R Console (64-bit)
File Edit Misc Packages Windows Help

> vss(x = vss.df[g2,6:35], n = 4, rotate = "oblimin",
+   fm = "mle", n.obs = length(g2))

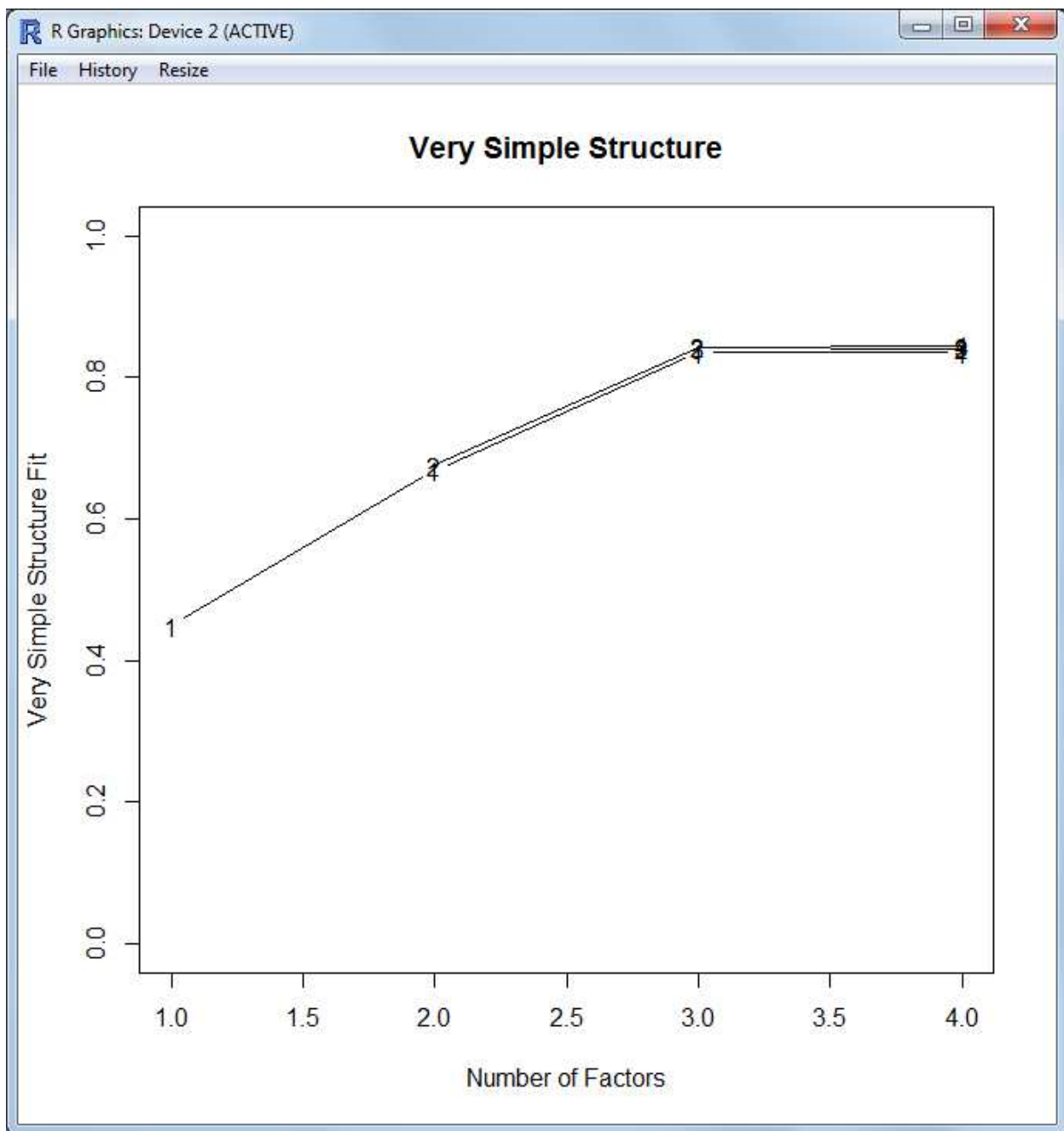
Very Simple Structure
Call: vss(x = vss.df[g2, 6:35], n = 4, rotate = "oblimin", fm = "mle",
  n.obs = length(g2))
VSS complexity 1 achieves a maximum of 0.84 with 3 factors
VSS complexity 2 achieves a maximum of 0.84 with 3 factors

The Velicer MAP achieves a minimum of 0 with 3 factors
BIC achieves a minimum of -2059.87 with 3 factors
Sample Size adjusted BIC achieves a minimum of -954.62 with 3 factors

Statistics by number of factors
  vss1 vss2   map dof chisq prob sqresid fit RMSEA   BIC SABIC complex eChisq SRMR eCRMS eBIC
1 0.45 0.00 0.0587 405 8098 0.00    54 0.45 0.140  5308  6594      1 32395 0.195 0.202 29604
2 0.67 0.68 0.0365 376 3757 0.00    32 0.68 0.096  1167  2361      1 13427 0.125 0.135 10837
3 0.84 0.84 0.0049 348  338 0.64    16 0.84 0.000 -2060  -955      1   195 0.015 0.017 -2203
4 0.84 0.84 0.0065 321  301 0.79    15 0.85 0.000 -1911  -891      1   174 0.014 0.017 -2038
> |

```

In this example all of the indices in the top table (“Very Simple Structure”) are in agreement; although both VSS complexity metrics display the same *maximum* for a 3–factor model and a 4–factor model. Looking at the first two columns of the “Statistics by number of factors” table shows the identical complexity *maximums* (0.84) for both the 3–factor model (row 3) and the 4–factor model (row 4) with both complexities 1 and 2 (columns 1 and 2). But, given the other indices agreement in support of the 3–factor model, that would be the model most appropriate. The plot (below) reinforces the interpretation of the tabular output above.



The plot (above) shows that the 3-factor model is meaningfully better than the 1-factor or 2-factor models and the 4-factor model does not show any improvement over the 3-factor model — which is evident because the number 4 in the plot is not [further] above the line associated with the 3-factor model (i.e. no gain or transition upward; as is the case from 1-factor to 2-factors and to 3-factors). Therefore, we fit the 3-factor model to our data using the ‘fa’ function (of the ‘psych’ package) and / or the ‘factanal’ function of the ‘stats’ package.

```

R Console (64-bit)
File Edit Misc Packages Windows Help

> fa(r = vss.df[g2,6:35], nfactors = 3, rotate = "oblimin", fm = "mle")
Factor Analysis using method = ml
Call: fa(r = vss.df[g2, 6:35], nfactors = 3, rotate = "oblimin", fm = "mle")
Standardized loadings (pattern matrix) based upon correlation matrix

```

	ML2	ML1	ML3	h2	u2	com
i1	-0.01	0.90	0.00	0.80	0.20	1
i2	-0.02	0.83	0.01	0.69	0.31	1
i3	0.04	0.78	0.01	0.63	0.37	1
i4	0.00	0.46	0.01	0.22	0.78	1
i5	0.03	0.68	0.00	0.47	0.53	1
i6	0.00	0.62	0.02	0.39	0.61	1
i7	0.00	0.59	-0.03	0.34	0.66	1
i8	0.01	0.46	0.02	0.22	0.78	1
i9	0.01	0.71	-0.01	0.50	0.50	1
i10	-0.04	0.69	0.00	0.46	0.54	1
i11	0.77	-0.03	0.03	0.59	0.41	1
i12	0.50	0.06	-0.03	0.26	0.74	1
i13	0.64	-0.01	-0.01	0.40	0.60	1
i14	0.72	0.01	0.02	0.53	0.47	1
i15	0.80	0.03	-0.01	0.64	0.36	1
i16	0.90	-0.01	-0.01	0.80	0.20	1
i17	0.79	0.00	-0.01	0.62	0.38	1
i18	0.79	0.02	0.01	0.64	0.36	1
i19	0.49	-0.06	-0.01	0.23	0.77	1
i20	0.67	-0.02	0.00	0.45	0.55	1
i21	0.04	0.00	0.60	0.37	0.63	1
i22	0.04	0.01	0.58	0.35	0.65	1
i23	0.02	0.04	0.52	0.29	0.71	1
i24	-0.01	0.00	0.71	0.50	0.50	1
i25	-0.01	-0.01	0.71	0.50	0.50	1
i26	-0.01	0.01	0.81	0.66	0.34	1
i27	-0.03	0.04	0.49	0.24	0.76	1
i28	0.02	0.02	0.64	0.42	0.58	1
i29	0.00	-0.02	0.72	0.51	0.49	1
i30	-0.02	-0.01	0.80	0.63	0.37	1

```

SS loadings          ML2  ML1  ML3
Proportion Var       0.17 0.16 0.15
Cumulative Var       0.17 0.33 0.48
Proportion Explained 0.36 0.33 0.31
Cumulative Proportion 0.36 0.69 1.00

With factor correlations of
      ML2  ML1  ML3
ML2  1.00 0.25 0.12
ML1  0.25 1.00 0.25
ML3  0.12 0.25 1.00

Mean item complexity = 1
Test of the hypothesis that 3 factors are sufficient.

The degrees of freedom for the null model are 435 and the objective function was 14.12 with ChiS

```

Note: the last few lines of output from the 'fa' function are cut off (i.e. not shown).

```

R Console (64-bit)
File Edit Misc Packages Windows Help

> factanal(vss.df[g2,6:35], factors = 3, rotation = "oblimin")

Call:
factanal(x = vss.df[g2, 6:35], factors = 3, rotation = "oblimin")

Uniquenesses:
  i1  i2  i3  i4  i5  i6  i7  i8  i9  i10 i11 i12 i13 i14 i15 i16
0.197 0.311 0.368 0.784 0.531 0.612 0.663 0.780 0.502 0.542 0.407 0.739 0.599 0.467 0.355 0.203
  i17 i18 i19 i20 i21 i22 i23 i24 i25 i26 i27 i28 i29 i30
0.378 0.365 0.774 0.553 0.631 0.651 0.713 0.504 0.504 0.338 0.756 0.584 0.485 0.365

Loadings:
  Factor1 Factor2 Factor3
i1          0.898
i2          0.833
i3          0.782
i4          0.464
i5          0.678
i6          0.618
i7          0.588
i8          0.462
i9          0.705
i10         0.687
i11  0.773
i12  0.497
i13  0.637
i14  0.725
i15  0.796
i16  0.896
i17  0.790
i18  0.791
i19  0.487
i20  0.673
i21          0.601
i22          0.582
i23          0.520
i24          0.706
i25          0.709
i26          0.813
i27          0.485
i28          0.638
i29          0.722
i30          0.802

  Factor1 Factor2 Factor3
SS loadings    5.163    4.711    4.442
Proportion Var  0.172    0.157    0.148
Cumulative Var  0.172    0.329    0.477

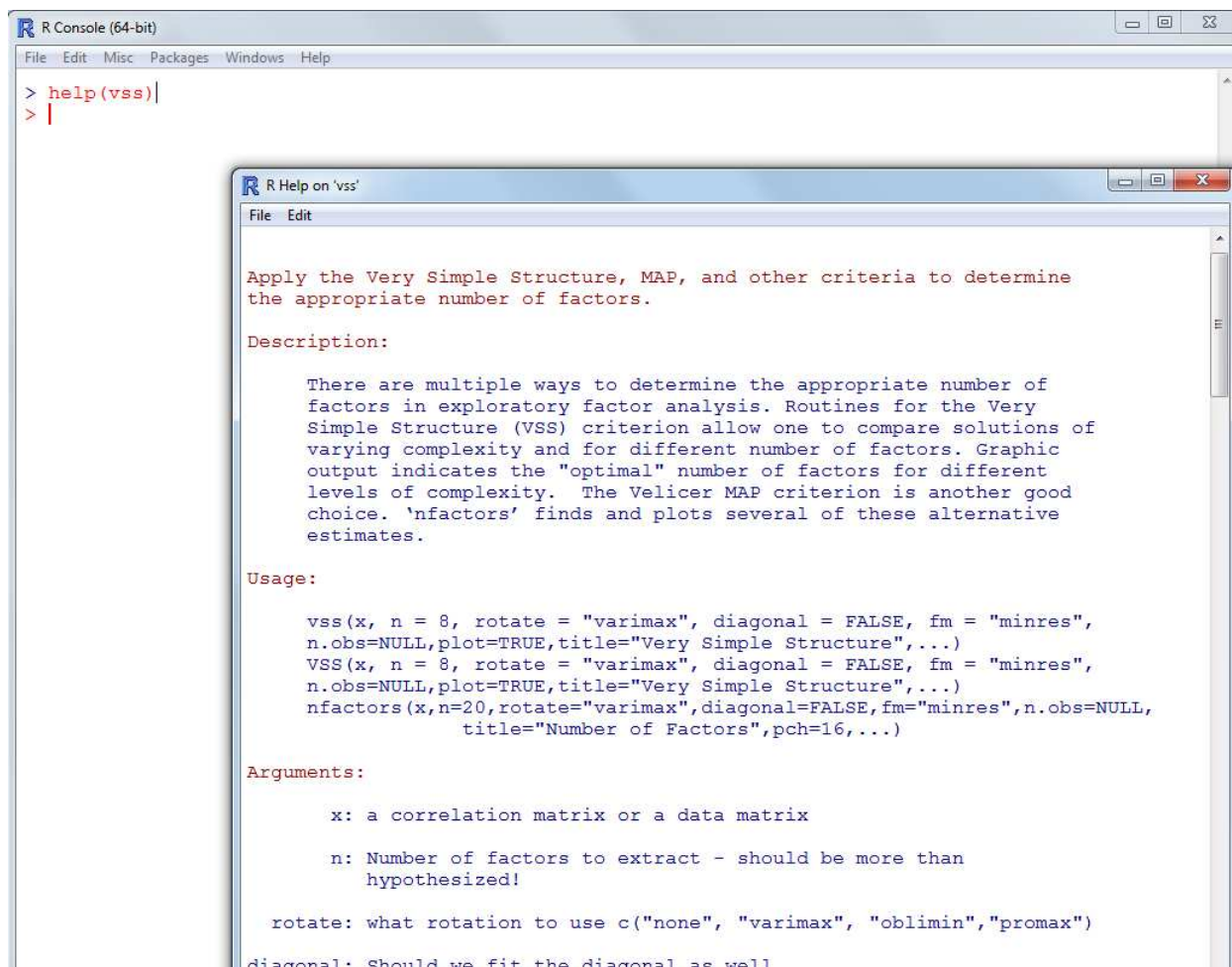
Factor Correlations:
  Factor1 Factor2 Factor3
Factor1   1.000    0.246   -0.252
Factor2   0.246    1.000   -0.119

```

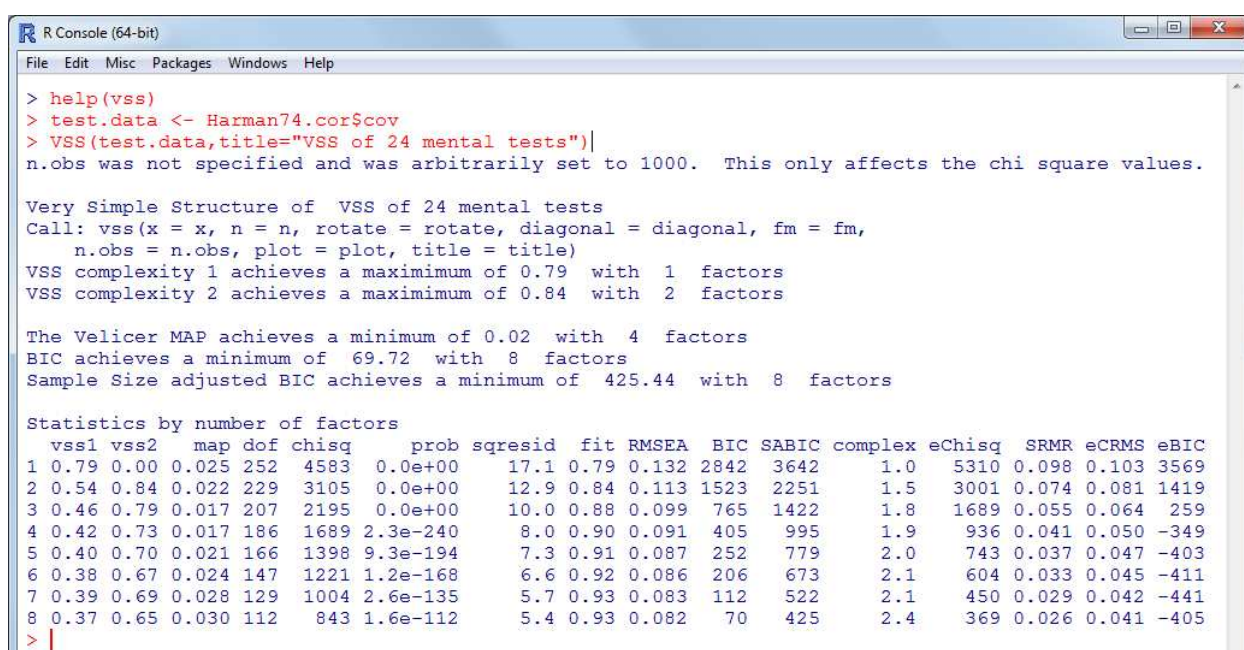
Note: last few lines of output from the ‘factanal’ function are cut off (i.e. not shown).

The next example is straight from the help file of the ‘vss’ function and is discussed here because it demonstrates a situation when the tables of output from the ‘vss’ function are not in agreement. When this situation occurs, one must rely upon the plot produced by the ‘vss’ function rather than the textual output. First, open the help file (here the plain text version is shown).

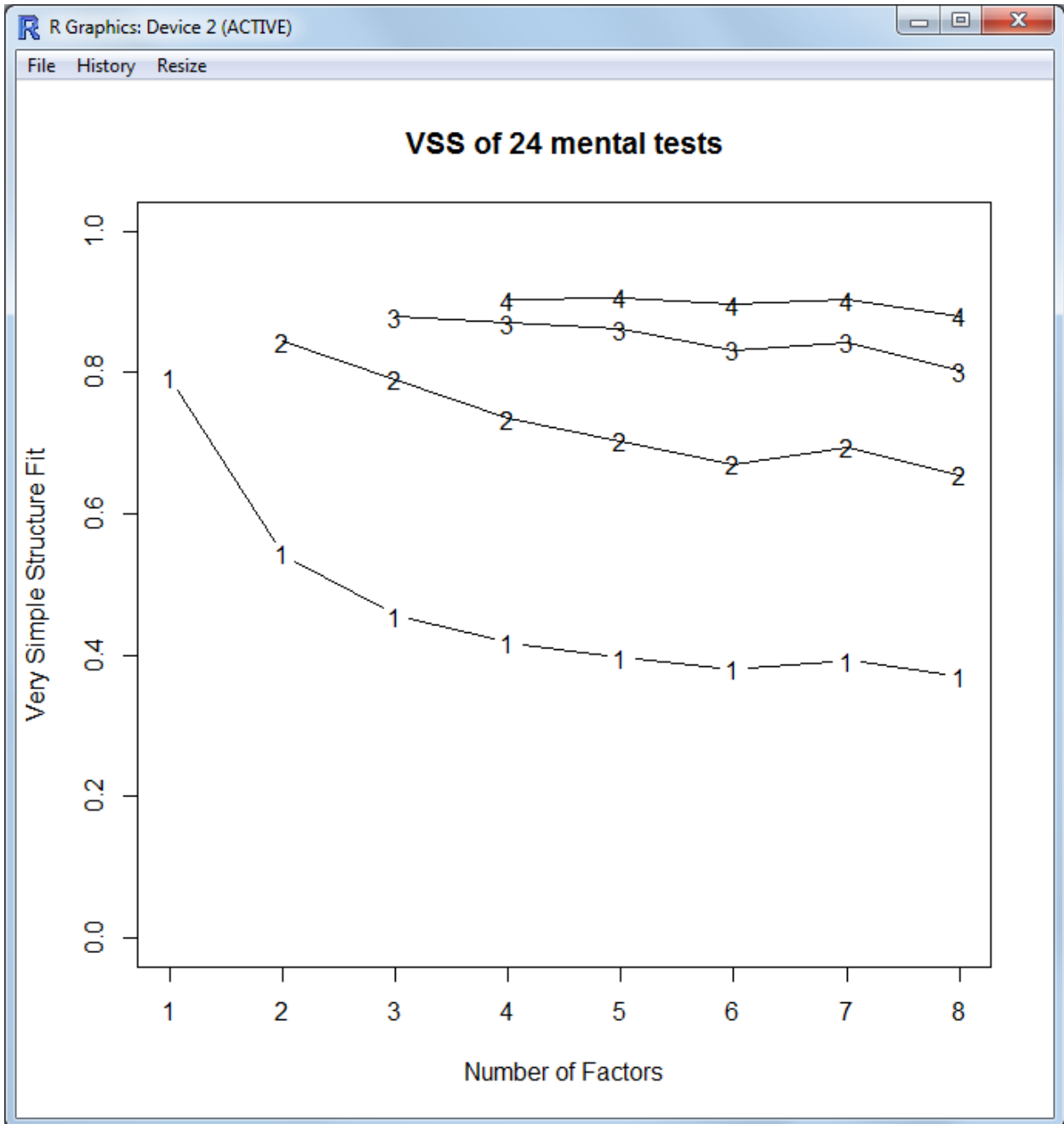




Next, scroll to the bottom of the help file and copy / paste the relevant lines of script into the R console.



As mentioned previously, the tables of statistics do not provide a clear answer to the question of which factor model is best (i.e. how many factors should be extracted). However, if we review the associated plot, we can clearly see the 4-factor model is the best (i.e. highest; even when embedded within models with more than 4 factors, with good separation from previous models).



## 2 Conclusions

The intent of this article was to raise awareness of the dangers of using only one criteria or method for deciding upon the number of factors to extract when conducting factor analysis. This article also demonstrated the ease with which an analyst can compute and evaluate several such criteria to reach a

more informed decision. More extensive examples of the data analysis solutions are available at the RSS *Do-it-yourself Introduction to R* course page:

[http://www.unt.edu/rss/class/Jon/R\\_SC/](http://www.unt.edu/rss/class/Jon/R_SC/)

Lastly, a copy of the script file used for the above examples is available here:

<http://www.unt.edu/rss/class/Jon/Benchmarks/VerySimpleStructure.R>

Until next time; remember what George Carlin said: “*just 'cause you got the monkey off your back doesn't mean the circus left town.*”

### 3 References & Resources

Bernaards, C., & Jennrich, R. (2014). The ‘GPArotation’ package. Documentation, including the package manual and the package vignette, available at CRAN:

<http://cran.r-project.org/web/packages/GPArotation/index.html>

Carlin, G. (1937 - 2008). *Just One-Liners*.

<http://www.just-one-liners.com/ppl/george-carlin>

Cattell, R. B. (1966). The scree test for the number of factors. *Multivariate Behavioral Research*, 1(2), 245 – 276.

Horn, J. (1965). A rationale and test for the number of factors in factor analysis. *Psychometrika*, 30(2), 179 – 185.

Horn, J. L., & Engstrom, R. (1979). Cattell’s scree test in relation to bartlett’s chi-square test and other observations on the number of factors problem. *Multivariate Behavioral Research*, 14(3), 283 – 300.

McDonald, R. P. (1999). *Test Theory: A Unified Treatment*. Mahwah, NJ: Erlbaum.

Pearson, K. (1901). On lines and planes of closest fit to systems of points in space. *Philosophical Magazine*, 2, 559 – 572.

Revelle, W. (2014). The ‘psych’ package. Documentation, including the package manual and the package vignette, available at CRAN:

<http://cran.r-project.org/web/packages/psych/index.html>

Revelle, W., & Rocklin, T. (1979). Very simple structure: An alternative procedure for estimating the optimal number of interpretable factors. *Multivariate Behavioral Research*, 14, 403 - 414. Available at: <http://personality-project.org/revelle/publications/vss.pdf>

Spearman, C. (1904). General Intelligence: Objectively Determined and Measured. *American Journal of Psychology*, 15, 201 – 292.

Statistics Canada. (2010). *Survey Methods and Practices*. Ottawa, Canada: Minister of Industry.

<http://www.statcan.gc.ca/bsolc/olc-cel/olc-cel?lang=eng&catno=12-587-X>

Thompson, B. (2004). *Exploratory and confirmatory factor analysis: Understanding concepts and applications*. Washington, DC: American Psychological Association.

Velicer, W. (1976). Determining the number of components from the matrix of partial correlations. *Psychometrika*, 41(3), 321 – 327.

This article was last updated on December 1, 2014.

This document was created using L<sup>A</sup>T<sub>E</sub>X