# Module 10.1: Simple (bivariate) Regression

#### Jon Starkweather, PhD jonathan.starkweather@unt.edu Consultant Research and Statistical Support

UNIVERSITY OF NORTH TEXAS Discover the power of ideas.

#### Introduction to Statistics for the Social Sciences



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- Introduction to Regression
- 3 NHST Example
  - Data
  - NHST Steps
  - Confidence Interval
  - Scatter Plots









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- NHST Steps
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- Scatter Plots





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- The following section offers a brief review of correlation.
- If anything regarding correlation is unclear, it is suggested you review the section on Measures of Relationship contained in Module 3: Describing Data.



### What is Correlation?

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- Typically, we will be using two continuous (or nearly so) variables; for example, depression scores, reaction time, age, heart rate, number of words or letters or symbols recalled, etc.
  - However, you can have categorical variables in correlation (e.g., point biserial correlation).



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Correlation Review Regression NHST 10.1 Summary

### How do we describe the relationship?



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- The stronger the relationship between the variables, the greater the absolute value of *r*



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Adjusted Correlation:

$$r_{adj} = \sqrt{1 - \frac{(1-r^2)(n-1)}{n-2}}$$





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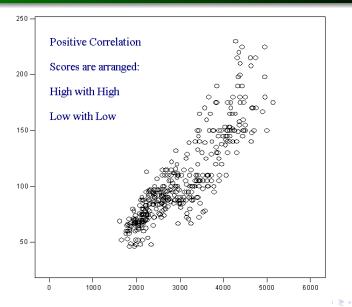
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- The statistical significance of a correlation is determined by comparing it to zero.

Correlation Review Regression NHST 10.1 Summary

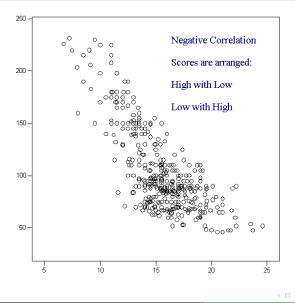
### **Positive Correlation**



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# **Negative Correlation**





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### **Correlation and Causation**



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#### **Correlation and Causation**

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- Three things may be able to explain the relationship:
  - X may cause Y
  - Y may cause X
  - Z (a third unknown variable) may be causing the relationship between X and Y.



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## Things that affect correlation



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• Restriction of Range.



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• Narrow distributions detract from the accuracy of r



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  - Outliers can *pull* a distribution's mean and bias correlation.



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Correlation Review Regression NHST 10.1 Summary

# Introduction to simple (bi-variate) Regression\*

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\*These slides were adapted with gracious permission from those produced by teaching and slide Guru, Dr. Mike Clark.



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Correlation Review Regression NHST 10.1 Summary

# From Correlation to Regression



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- In our scatter plots; the variable on the X-axis (the horizontal axis in Cartesian plane space) is called the *Predictor*
- The variable on the Y-axis (the vertical axis) is called the *Outcome*.



Correlation Review Regression NHST 10.1 Summary

## The formula for a Straight line



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  - X is the value(s) of the variable on the horizontal axis (X-axis).
- Once this line is specified, we can calculate the corresponding value of Y for any new value of X.

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## The Line of Best Fit, the Linear Prediction Rule

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 $\widehat{Y} = bX + a$ 

• Where  $\widehat{Y}$  is the *predicted* value of *Y*.

#### Least Squares Modeling

• When the relations between variables are expressed in this manner, we call the relevant equations mathematical *models*.



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# Least Squares Modeling

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- The intercept and coefficient are called *parameters* of a model.
- We assume that our models are causal models, such that the variable on the left-hand side of the equation is being caused by the variable(s) on the right side (not to be confused with establishing causality; X still does not necessarily cause Y).



# Terminology

When the values of Y in these models are called predicted values (sometimes abbreviated as Y-hat), they are given the symbol Ŷ.



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# Terminology

- When the values of Y in these models are called predicted values (sometimes abbreviated as Y-hat), they are given the symbol Ŷ.
- They are the values of Y that are implied or predicted by the specific parameters of the model and the values of X.



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- There are two important issues we need to deal with however:
  - Is the basic model correct (regardless of the value of the parameters)? That is, is a linear, as opposed to a quadratic/curvilinear, model the appropriate model for characterizing the relationship between two variables?
  - If the model is correct, what are the most correct parameter values for the model?



• The process of obtaining the correct parameters (assuming we are working with the right model) is called *Parameter Estimation*.



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- Often, theories specify the *form* of the relationship rather than the specific values of the parameters.
- The parameters themselves, assuming the basic model is correct, are typically estimated from the data. We refer to the estimation processes as *calibrating the model*.
- We need a method for choosing parameter values which will give us the best representation of the data points.

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## Simple Parameter Estimation example data

• We collect scores from 4 participants on two variables.



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Correlation Review Regression NHST 10.1 Summary

#### Simple Parameter Estimation example data

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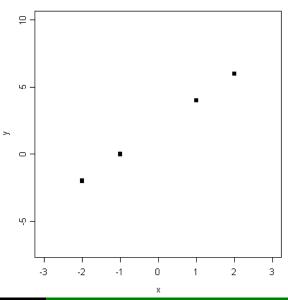


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  - (-2, -2), (-1, 0), (1, 4), (2, 6)

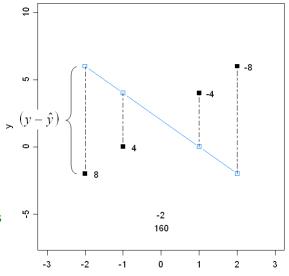


- Assuming we believe there is a linear relationship between x and y.
- Which set of parameter values will bring us closest to representing the data accurately?



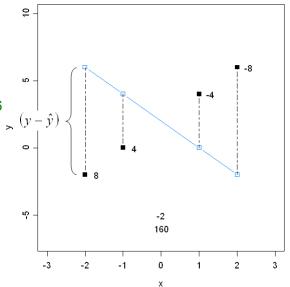
Starkweather Module 10.1

- Model  $\hat{y} = 2 2x$  in light blue
- Pick some parameter values and see how well the model does.
- Quantify "how well" with the difference between the model's predicted values (ŷ) and the actual values (y)
- This difference,  $(y \hat{y})$  is called *error in prediction* or *residual*.

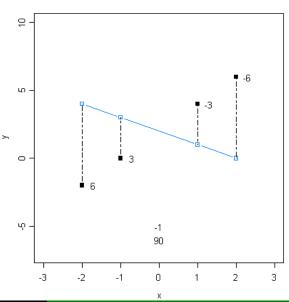


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- Model  $\hat{y} = 2 2x$  in light blue
- So, for the first data point, x = -2
- The model predicts  $\hat{y} = 6$ because:  $\hat{y} = 2 - 2(-2) = (y - \hat{y})$
- The residual  $(y \hat{y}) = -2 6 = -8$ .

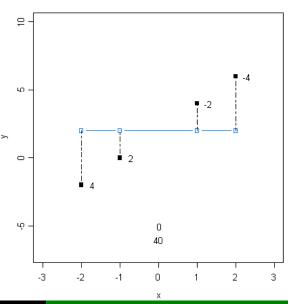


- $\hat{y} = 2 1x$
- Try a different value for *b* and see what happens.
- The predicted values are getting closer, but still off quite a bit.



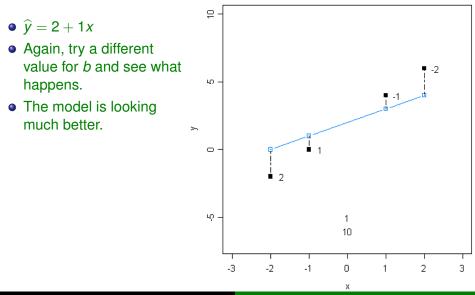


- Again, try a different value for *b* and see what happens.
- The model is getting better (smaller residuals), but can still be improved.



Starkweather Moc

Module 10.1

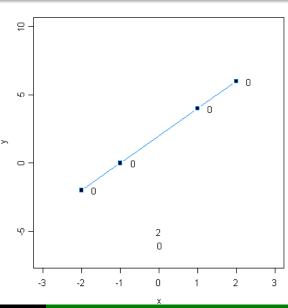


Starkweather Mod

Module 10.1



- Again, try a different value for *b* and see what happens.
- Perfect.
- Zero residuals!
- Of course, this never happens with real data.
- There will always be *some* residual.



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- The process of finding this minimum value is called **Least-squares Estimation**.
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- This is why you will often hear researchers refer to regression as Ordinary Least-Squares (OLS) regression.

# Estimate b

#### • Estimating the Slope (the regression coefficient)



3

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## Estimating a

#### • Estimating the Y-intercept: $a = \overline{Y} - b\overline{X}$



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- Where the means are based on the sets of Y and X data values and *b* is the slope.
- These calculations ensure that the regression line passes through the point on the scatterplot defined by the two means.



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- Remember, Z-scores are also called 'standard scores', so we would have a *standardized* regression coefficient or beta coefficient.



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- In simple (bi-variate) regression, r = β meaning; the correlation between the predictor variable and the outcome variable equals the standardized regression coefficient (β).



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  - Reflecting how well our model (with its parameters) *fits* the data.

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### Interpreting a Regression Summary

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• Same as  $\eta^2$  in ANOVA.

#### • Determining if the Regression Model is significant.



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- We have determined the form of the relationship (Y = aX + b) and its strength (*r* or  $r_{adj}$ ), as well as the Model's effect size (variance of the outcome accounted for by the predictor using  $r^2$  or  $r_{adj}^2$ ).



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- Determining if the Regression Model is significant.
- We have determined the form of the relationship (Y = aX + b) and its strength (*r* or  $r_{adj}$ ), as well as the Model's effect size (variance of the outcome accounted for by the predictor using  $r^2$  or  $r_{adj}^2$ ).
- But, does a prediction based on this model do a better job than just predicting the mean of Y for any new value of X?
  - After all; if Y is normally distributed, then  $\overline{Y}$  is our best guess for an unknown score on it.
- ANOVA is used to answer that question.



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# Sums of (regression) Squares

 We can calculate an ANOVA for testing whether or not r<sup>2</sup> is significantly different from 0 using the different partitions of variance discussed above.



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$$SOS_Y = \sum \left(Y - \overline{Y}\right)^2$$

### **Regression (ANOVA) Summary Table**

Source	SOS	df	MS	F <sub>calc</sub>
Predicted	$SOS_{\widehat{Y}} = \sum \left(\widehat{Y} - \overline{Y}\right)^2$	1	$\frac{SOS_{\widehat{Y}}}{df_{\widehat{Y}}}$	$rac{MS_{\widehat{Y}}}{MS_e}$
Error	$SOS_e = \sum \left(Y - \widehat{Y}\right)^2$	n – 2	<u>SOSe</u> dfe	
Total	$SOS_Y = \sum \left(Y - \overline{Y}\right)^2$	<i>n</i> – 1		



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$$t = \frac{b-b*}{S_b} = \frac{b-0}{\frac{S_{Y,X}}{S_{X}*\sqrt{n-1}}} = \frac{(b)(S_X)(\sqrt{n-1})}{S_{Y,X}}$$

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  - Outliers can *pull* a distribution's mean and bias correlation.



Data Steps CI Scatter Plots

#### Example Data and Preliminary Calculations

 Students (n = 10) were randomly sampled, then their Stress (X) and Achievement (Y) levels were recorded.



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• Students (n = 10) were randomly sampled, then their Stress (X) and Achievement (Y) levels were recorded. Stress (X) Achievement (Y)  $\sum X = 1400.53$  $\sum Y = 4838.48$  $\overline{X} = 140.053$  $\overline{Y} = 483\ 848$  $S_{X} = 27.866$  $S_V = 39.878$  $S_Y^2 = 776.541$  $S_V^2 = 1590.255$  $\sum \left(X - \overline{X}\right) \left(Y - \overline{Y}\right) = 5145.125$ 

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- Data on the next slide.

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code	X	$X - \overline{X}$	Y	$Y - \overline{Y}$	$\left( X-\overline{X} ight) \left( Y-\overline{Y} ight)$
16	151.53	11.477	475.92	-7.928	-90.990
20	135.97	-4.083	510.29	26.442	-107.963
28	206.71	66.657	568.19	84.342	5621.985
33	107.38	-32.673	485.65	1.802	-58.877
41	136.99	-3.063	493.93	10.082	-30.881
74	145.29	5.237	485.00	1.152	6.033
90	142.58	2.527	437.22	-46.628	-117.829
93	125.51	-14.543	444.78	-39.068	568.166
95	107.29	-32.763	501.72	17.872	-585.540
97	141.28	1.227	435.78	-48.068	-58.979
	1400.53		4838.48		5145.125
					RSS Research and Statistical Support

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- Population 1: UNT students' Stress levels (X).
   Population 2: UNT students' Achievement levels (Y).
- Hypothesis 1: The shared variance between X and Y will be significantly greater than zero.





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Step 2



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http://www.math.unb.ca/~knight/utility/t-table.htm





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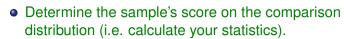
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Step 4

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• Which gives us the following model:

$$\hat{Y} = 380.741 + 0.7362X$$

## **Regression (ANOVA) Summary Table**

Recall the Regression (ANOVA) Summary Table and notice we needed the model (parameters) to get the  $\hat{Y}$  values, which are needed to calculate two *SOS* 

Source	SOS	df	MS	F <sub>calc</sub>
Predicted	$SOS_{\widehat{Y}} = \sum \left(\widehat{Y} - \overline{Y}\right)^2$	1	$\frac{\textit{SOS}_{\widehat{Y}}}{\textit{df}_{\widehat{Y}}}$	$rac{MS_{\widehat{Y}}}{MS_e}$
Error	$SOS_{e} = \sum \left(Y - \widehat{Y}\right)^{2}$	n – 2	SOS <sub>e</sub> df <sub>e</sub>	
Total	$SOS_Y = \sum (Y - \overline{Y})^2$	<i>n</i> – 1		
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## Calculating *Predicted* Sums of Squares $(SOS_{\hat{v}})$

$\hat{\mathbf{v}}$	$\overline{\mathbf{v}}$	~ _	(
r	Ŷ	$\widehat{Y} - \overline{Y}$	$(\overline{Y} - \overline{Y})^{-}$
492.2975 4	83.848	8.449	71.393
480.8421 4	83.848	-3.006	9.036
532.9214 4	83.848	49.073	2408.199
459.7939 4	83.848	-24.054	578.601
481.5930 4	83.848	-2.255	5.085
487.7035 4	83.848	3.856	14.865
485.7084 4	83.848	1.860	3.461
473.1413 4	83.848	-10.707	114.633
459.7276 4	83.848	-24.120	581.793
484.7513 4	83.848	0.903	0.816

$$SOS_{\widehat{Y}} = \sum \left(\widehat{Y} - \overline{Y}\right)^2 = 3787.881$$



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## Calculating *Error* Sums of Squares (*SOS*<sub>e</sub>)

Y	Ŷ	$Y - \widehat{Y}$	$\left(Y-\widehat{Y}\right)^2$
475.92	492.2975	-16.377	268.221
510.29	480.8421	29.448	867.181
568.19	532.9214	35.269	1243.874
485.65	459.7939	25.856	668.539
493.93	481.5930	12.337	152.202
485.00	487.7035	-2.704	7.309
437.22	485.7084	-48.488	2351.125
444.78	473.1413	-28.361	804.365
501.72	459.7276	41.992	1763.360
435.78	484.7513	-48.971	2398.191
			$\downarrow$

$$SOS_e = \sum \left(Y - \widehat{Y}\right)^2 = 10524.37$$



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## Calculating *Total* Sums of Squares $(SOS_Y)$

Y	Ŷ	$Y - \overline{Y}$	$\left(Y-\overline{Y}\right)^2$
475.92	483.848	-7.928	62.853
510.29	483.848	26.442	699.179
568.19	483.848	84.342	7113.573
485.65	483.848	1.802	3.247
493.93	483.848	10.082	101.647
485.00	483.848	1.152	1.327
437.22	483.848	-46.628	2174.170
444.78	483.848	-39.068	1526.309
501.72	483.848	17.872	319.408
435.78	483.848	-48.068	2310.533
			↓

$$SOS_{Y} = \sum \left(Y - \overline{Y}\right)^{2} = 14214.25$$



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## **Regression (ANOVA) Summary Table**

• Finally, we can construct the Summary Table with the correct values.

Source	SOS	df	MS	F <sub>calc</sub>
Predicted	3787.881	1	3787.881	2.879
Error	10524.370	8	1315.546	
Total	14312.250	9		



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  - It really depends upon the context of the particular study (i.e. previous research findings with these variables).

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- So, you can see the significance test of b is really unnecessary when r<sup>2</sup> is not significantly different from zero.

#### Calculating the t test for b

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Data Steps CI Scatter Plots

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$$t = \frac{b - b*}{S_b} = \frac{0.7362 - 0}{0.4339} = 1.6967$$

Data Steps CI Scatter Plots

## Step 5: Hypothesis 2

• Compare and make a decision.



## Step 5: Hypothesis 2

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- No surprises here; since t<sub>calc</sub> = 1.6967 < 1.860 = t<sub>crit</sub> we fail to reject the null hypothesis and conclude that our sample does not provide evidence to support the idea that Stress levels (X) are a significant predictor of Achievement (Y).



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- Of course, we can still calculate a confidence interval for *b* which will include zero indicating a lack of statistical significance.



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which then become...

$$LL = -1.860 * .4339 + .7362 = -0.0708$$
  
 $UL = +1.860 * .4339 + .7362 = 1.5432$ 



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- So, our LL = -0.0708 and our UL = 1.5432; which means our interval includes zero.
  - Like the NHST we can conclude that our *b* is not significantly different from zero.
- We interpret this CI as; if we drew an infinite number of samples of UNT students and measured their Achievement and Stress levels, 95% of the regression coefficients (*b*) would be between -0.708 and 1.5432.



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  - Or, if a gap exists between zero and the location of the first data point; then the axis with the gap instead has a jagged segment between the origin and the first data point.
  - The jagged segment indicates distance along the axis between zero and the first *tick mark* or number of the scale (of that axis).

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    - This line would represent a *positive* correlation (high scores on X tend to have high scores on Y *and* low scores on X tend to have low scores on Y).

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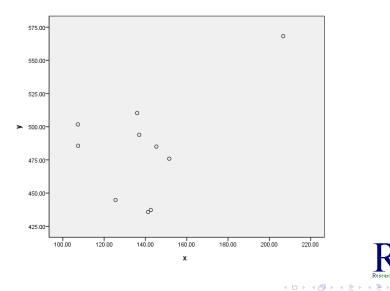


Data Steps CI Scatter Plots

Research and Statistical Support

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### **Basic Scatter Plot**



• The previous scatter plot was produced with SPSS; the following scatter plot was produced with R.



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  - Boxplots for each axis allow us to see how the data is distributed along each individual axis.



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Data Steps CI Scatter Plots

### Grid lines and Boxplots

Faint Grid lines



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• Again, imagine a line from the lower left to the upper right when looking at the plot.



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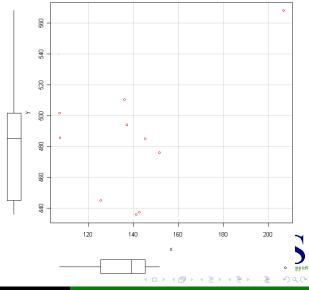
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Correlation Review Regression NHST 10.1 Summary

ta Steps CI Scatter Plots

### Scatter Plot w/grid lines and box plots on each axis

Notice the outlier; also shown at the bottom for the x-axis boxplot.



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Data Steps CI Scatter Plots

### The Regression Line

 $\widehat{Y} = 380.741 + .7362X$ 



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  - Notice the y-intercept (a = 380.741) does not appear correct because, the scale of each axis does not originate with zero; if they did, the regression line would intersect the y-axis at 380.741.



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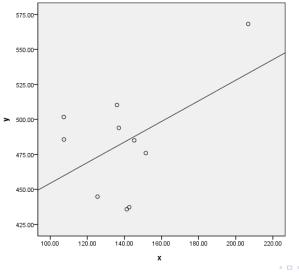


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  - Looking at the following scatter plot, we can also see the slope (b = .7362) is fairly steep.
    - Slope = rise over run in decimal form.
  - Recall, the best fit regression line represents the points which would be predicted  $(\widehat{Y})$  by our model for Y, given new values of X.

### Scatter Plot w/regression line of best fit







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RSS Research and Statistical Support

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  - Complete results from SPSS are below, then the scatter SSS plot follows.

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### Population Regression SPSS Output

### Notice, in Bi-variate regression; correlation equals beta $(r = \beta)$

#### Model Summary<sup>b</sup>

Model	R	R Square	Adjusted R Square	Std. Error of the Estimate			
1	.547ª	.300	.293	30.58640			
a. Predictors: (Constant), x							

b. Dependent Variable: y

ANOVA<sup>b</sup>

	Model		Sum of Squares df		Mean Square	F	Sig.
Γ	1	Regression	39232.697	1	39232.697	41.936	.000ª
I		Residual	91681.741	98	935.528		
l		Total	130914.437	99			

a. Predictors: (Constant), x

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Coefficients<sup>a</sup>

# Also, in SPSS: R is used for r and B is used for b

Unstandardized Coefficie		d Coefficients	Standardized Coefficients			95.0% Confidence Interval for B		
Model		В	Std. Error	Beta	t	Sig.	Lower Bound	Upper Bound
1	(Constant)	366.458	18.156		20.184	.000	330.428	402.489
	х	.779	.120	.547	6.476	.000	.540	1.017

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a. Dependent Variable: y



Correlation Review Regression NHST 10.1 Summary Scatter Plots

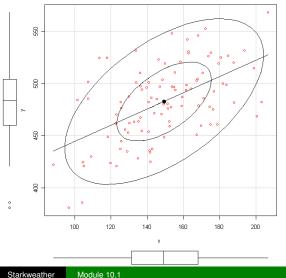
### Population (n = 100) Scatter Plot

 $\widehat{Y} = 366.458 + .779X$ 

Notice the ellipses enclose the bulk (60% and 90%) of the data.

Also, the solid dot is the *centroid*; which is the point at  $(\overline{X}, \overline{Y})$ 

Also, 2 outliers on Y (y-axis boxplot).



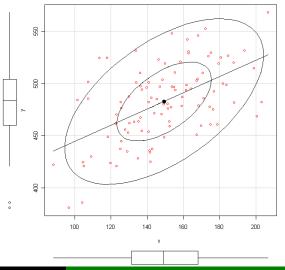
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Correlation Review Regression NHST 10.1 Summary Data Steps CI Scatter Plots

### Population (n = 100) Scatter Plot

 $\widehat{Y} = 366.458 + .779X$ 

A *new* participant has a Stress level (X) of 100; what do you predict he or she will score on Achievement?



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Correlation Review Regression NHST 10.1 Summary Data Steps CI Scatter Plots

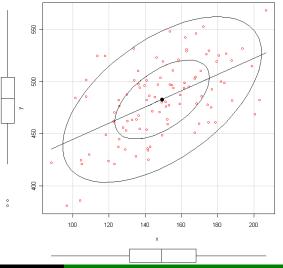
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 $\hat{Y} = 444.358$ 

444.358 = 366.458 + .779(100)



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  - WAIS and HALE: r = 1.00



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  - HALE can be described as normally distributed in the population of the U.S. with:  $\mu = 50, \sigma = 10$
  - WAIS and HALE: *r* = 1.00
- QUESTION: If U.S. citizen John Doe scores a 130 on the WAIS, what would be a good guess for his HALE?

### • Think about it for a few minutes.



- Think about it for a few minutes.
- Think about what a scatter plot would look like, just given the information from the previous slide.



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- Think about it some more ;)



### ANSWER

 Recall, in bi-variate regression; r = β and beta is the standardized regression coefficient, or standardized slope (i.e. rise over run when the variables are standardized).



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  - John Doe would likely have a 70 on the HALE.

- Recall, in bi-variate regression; r = β and beta is the standardized regression coefficient, or standardized slope (i.e. rise over run when the variables are standardized).
  - The generic standardized regression equation is:  $Z_Y = \beta * Z_X$
- So, if both WAIS and HALE are transformed into standardized scores, then they both would have a mean of zero and a standard deviation of 1.00.
- Then, since the correlation is a perfect 1.00, we can conclude that a standardized score of 2 on WAIS (x-axis) corresponds to a standardized score of 2 on HALE (y-axis).
  - John Doe would likely have a 70 on the HALE.
  - Scatter Plots follow.



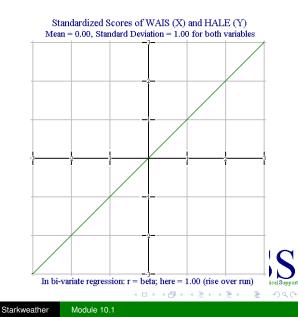
Correlation Review Regression NHST 10.1 Summary

Data Steps CI Scatter Plots

#### Standardized Scatter Plot: Centroid is (0, 0)

 $Z_{\widehat{Y}} = 1.00 Z_X$ 

The key is that the 'tick marks' (numbers) are **at** each standard deviation and r = 1.00.

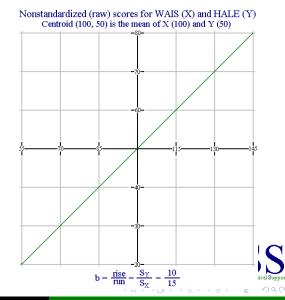


#### **Unstandardized Scatter Plot**

 $\widehat{Y} = 50 + \frac{10}{15}X$ 

The key is that the 'tick marks' (numbers) are **at** each standard deviation and r = 1.00.

John Doe: X = 130,  $\widehat{Y} = 70$ 



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  - Using  $r_{adj}^2$  as an unbiased estimate of variance accounted for effect size.



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  - Confidence interval for b (Cl<sub>95</sub>)
  - Scatter Plots



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### This concludes Module 10.1

A firm understanding of the topics covered here and previously will be necessary for understanding future topics.

- Next time Module 11.
- Next time we'll be covering Categorical data analysis techniques.
- Until next time; have a nice day.

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