Module 11: Nominal and Ordinal Variable Analysis

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Introduction to Statistics for the Social Sciences



Starkweather Module 11

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http://www.unt.edu/rss/Instructional.htm





Chi-square test

- One-way Classification Tables
- Multi-way Contingency Tables
- Effect Size
- Kappa





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- 2 Wilcoxon's Ranks tests
 - Wilcoxon's Rank-Sum Test
 - Wilcoxon's Matched-Pairs Signed-Ranks Test



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- 4 Summary of Module 11



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 - Assumptions about population distributions.



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 - Concerned with population values (i.e. parameters).
 - Require Interval and/or ratio scaled variables.
 - Assumptions about population distributions.
- This module (11) concerns itself with *Nonparametric* statistics.



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 - Few if any assumptions.
 - Sometimes called distribution-free tests because, they do not make assumptions about a population distribution.
- Unfortunately, nonparametric tests tend to have less power or sensitivity to detect significance than their parametric partners.



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- Both Chi-square tests use the same formula and are based on the distribution of Chi-square.



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 - Symbol: χ^2



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• Notice in the table linked above, the sixth column corresponds to a significance level of 0.05 where:



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 - The first column is degrees of freedom (df)



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 - E is the frequency expected if the null hypothesis were true.

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One-way Classification Table Example

• The One-way Chi-square test is the Goodness-of-fit test.



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Chi-square Wilcoxon Kruskal-Wallis Summary One-way Two-way Effect Size kappa

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 - The alternative hypothesis: $H_1 : E \neq O$
- Instead, we found: 32 Freshmen, 28 Sophomores, 23 Juniors, and 17 Seniors.



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 - The null hypothesis would be: $H_0: E = O$
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- Instead, we found: 32 Freshmen, 28 Sophomores, 23 Juniors, and 17 Seniors.
- This study design constitutes a one-way classification table because, there is only one variable (class level) with RSS multiple categories.

The One-way Classification Table

	Freshmen	Sophomore	Junior	Senior
Observed	32	28	23	17
Expected	25	25	25	25



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	Freshmen	Sophomore	Junior	Senior
Observed	32	28	23	17
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• Degrees of Freedom (*df*) is the number of Categories or Columns minus 1.



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Observed	32	28	23	17
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- df = C 1 = 4 1 = 3

http://www.medcalc.be/manual/chi-square-table.php

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Calculate Chi-square

• Using the formula from above,



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$$\chi^2 = \sum \frac{(O-E)^2}{E}$$



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$$\chi^2 = \frac{(32-25)^2}{25} + \frac{(28-25)^2}{25} + \frac{(23-25)^2}{25} + \frac{(17-25)^2}{25} = 5.04$$



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Calculate Chi-square

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• And since $\chi^2_{calc} = 5.04 < 7.815 = \chi^2_{crit}$ we fail to reject the null hypothesis and conclude that this sample does not indicate a significant difference between the observed and expected frequencies of class level.

Multi-way Chi-square

• When we have more than one categorical variable, we call the chi-square test a test of *Independence*.



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 - Are the cells of the table *Independent* of one another, or is there some relationship occurring among them.
- In the one-way example above, we called the table a *classification table* because we were classifying frequencies on one variable.
- In the multi-way situation, we call the table a *contingency table* because, the frequencies of one variable are contingent upon another (or more than one) variable.



A Two-way Example

• Suppose we wondered about the gender frequency of students entering the UNT Administration building from above?



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- A 2 X 4 design (Gender by Class Level).



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- Suppose we wondered about the gender frequency of students entering the UNT Administration building from above?
- A 2 X 4 design (Gender by Class Level).

		Class	Level		
Gender	Freshmen	Sophomore	Junior	Senior	Total
Male	32	28	23	17	100
Female	28	29	20	15	92
Total	60	57	43	32	192



• In the one-way design, expected frequencies were simply even proportions; but here, with a more complex design, we must calculate the expected frequencies which are *contingent* upon two variables.



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• Where *E_{ij}* is a particular cell, *R_i* is the row total, *C_j* is the column total, and *n_t* is the total number of individuals (or cases).

Expected Frequencies for the current example

• For the current example, we have the following Expected frequencies for each cell:



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Chi-square Wilcoxon Kruskal-Wallis Summary

One-way Two-way Effect Size kappa

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Which leads to:

 $E_{11} = 31.25$ $E_{12} = 29.69$ $E_{13} = 22.40$ $E_{14} = 16.67$

 $E_{21} = 28.75$ $E_{22} = 27.32$ $E_{23} = 20.60$ $E_{24} = 15.33$

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Table with Expected Frequencies

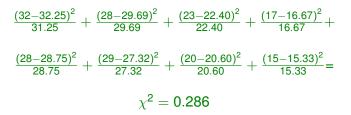
• Here we have the Expected Frequencies for each cell, listed in parentheses.

		Class	Level		
Gender	Freshmen	Sophomore	Junior	Senior	Total
Male	32(31.25)	28(29.69)	23(22.40)	17(16.67)	100
Female	28(28.75)	29(27.32)	20(20.60)	15(15.33)	92
Total	60	57	43	32	192

 Of course, you can not have 31.25 persons (frequencies), so you could round to the nearest whole number.

Calculating χ^2 for the two-way example

• Recall the formula for $\chi^2 = \sum \frac{(O-E)^2}{E} =$





• Recall, earlier we said df = C - 1



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- So, *df* = (*R* − 1)(*C* − 1)



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 - Where R = the number of rows and C = the number of columns.



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• So, our
$$\chi^2_{crit} = 7.815$$
 is the same.



• So our $\chi^2_{calc} = 0.286 < 7.185 = \chi^2_{crit}$ we fail to reject the null hypothesis and we conclude that there was not a relationship between Gender and Class Level.



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Two-way Example Results

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- Stated still another way, the Observed frequencies for each cell did not differ significantly from the Expected frequencies.
- Like with correlation, chi-square is very sensitive to sample size.
 - If given a large enough sample, any chi-square analysis will be significant.

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 - To answer that question, simply divide the number of Freshmen by the number of *not Freshmen* for the Male row.
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- Odds of a male also being a Freshman: $\frac{32}{68} = 0.4706$ or nearly 50/50 odds.
- Stated another way: there is a 47.06% chance a male entering the building is also a Freshman.

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Phi as 2 X 2 Contingency Effect Size

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- Of course, it is limited to the 2 X 2 situation only.



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- Cohen's kappa corrects this deficiency.

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Agreement Data

NE = Not Effective, E = Effective, HE = Highly Effective.

		Faculty 1		
Faculty 2	NE	E	HE	Total
NE	4	0	0	4
E	0	5	1	6
HE	0	3	15	18
Total	4	8	16	28



kappa

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 - So, the probability of both faculty agreeing on 'Effective' for one student is .2857*.2143 = .0612.



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 - So, the probability of both faculty agreeing on 'Effective' for one student is .2857*.2143 = .0612.
 - Which is not a lot, but across all 28 students, we can expect .0612*28 = 1.71 agreements just by random chance.



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• Where f_o is the observed frequencies on the diagonal and f_e is the expected frequencies on the diagonal.



Calculating the Expected Frequencies



Calculating the Expected Frequencies

$$E_{ij} = \frac{R_i C_j}{n_t}$$



• Use the same formula from earlier to calculate the Expected Frequencies:

 $F_{ii} = \frac{R_i C_j}{C_j}$

• For Not Effective (NE):
$$(4^*4)/28 = .571$$



$$E_{ij} = \frac{R_i C_j}{n_t}$$

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- For Highly Effective (HE): (18*16)/28 = 10.286



$$E_{ij} = \frac{R_i C_j}{n_t}$$

- For Not Effective (NE): (4*4)/28 = .571
- For Effective (E): (6*8)/28 = 1.714
- For Highly Effective (HE): (18*16)/28 = 10.286
- Then, sum them to get $f_e = 12.571$



Calculating the Observed Frequencies

• Simply add up the observed frequencies to get fo



Calculating the Observed Frequencies

- Simply add up the observed frequencies to get f_o 4+5+15=24
- Now we can calculate kappa.



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• So, agreement is really lower than the 85.71% from above; after accounting for chance it is 74.07%.



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RSS Research and Statistical Support

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- However, as mentioned previously, non-parametric tests tend to have less power than their parametric companions.
 - The Rank-Sum test has less power than the Independent Samples *t* test.
- The general idea of the Rank-Sum test is to test whether two samples originated with the same population, similar to the Independent Samples *t* test.
 - However, it is not specifically tied to *mean* differences, but rather; differences in central tendency.

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http://www.unt.edu/rss/class/Jon/ISSS_SC/Module011/ws_tables/

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 - One-tailed test: police officers < taxi drivers.
- To test this we will first rank all the scores.



Ranked Data

	Raw Scores	Rank
		naiin
Police	8	1
Officers	15	5
	12	3
	10	2
	13	4
Taxi	27	9
Drivers	28	10.5
	19	7
	17	6
	26	8
	28	10.5

Tied scores get tied ranks half-way between the two whole number ranks they would occupy if sequential.



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• Since our calculated $W_s = 15 < 20 = W_s$ critical value; we reject the null hypothesis and conclude that the two groups *are* significantly different.

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- Notice, the table provides $2\overline{W}$ in the right most column.
- Then, if W'_s is larger than the critical value, we would reject the null and conclude that the taxi drivers scored significantly higher on the Driving Anger scale.

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Normal Approximation

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 - And the standard deviation of the distribution of W_s is: $\sqrt{\frac{n_1n_2(n_1+n_2+1)}{12}}$
- So, the *z* score is calculated using:

$$Z = \frac{\text{statistic} - \text{mean}}{\text{standard deviation}} = \frac{W_s - \frac{n_1(n_1 + n_2 + 1)}{2}}{\sqrt{\frac{n_1 n_2(n_1 + n_2 + 1)}{12}}}$$

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 So we could say the police officers scored significantly lower than the taxi drivers because a critical z value of -1.64 corresponds to a one-tailed test of z at 0.05 (negative because we hypothesized the police would be **lower**).

http://www.mathsisfun.com/data/standard-normal-distribution-table.html



Wilcoxon's Matched-Pairs Signed-Ranks test

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- Like the previous Wilcoxon test, this one works with ranks and the sum of ranks.



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- We would expect the post test scores to be lower than the pretest scores.
 - One-tailed test, lower end



Example Data

Pre	Post	Difference	Rank of difference	Signed Rank
15	8	7	2.5	2.5
18	10	8	4.5	4.5
17	8	9	6.5	6.5
19	11	8	4.5	4.5
20	13	7	2.5	2.5
22	12	10	8.5	8.5
16	18	-2	1	-1
24	12	12	10	10
23	14	9	6.5	6.5
21	11	10	8.5	8.5

 $T + = \sum positive ranks = 54$

 $T - = \sum negative \ ranks = -1$

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http://comp9.psych.cornell.edu/Darlington/wilcoxon/wilcox5.htm

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T_{calc} versus T_{crit}

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- So, the T_{crit} (labeled S in the table linked above), for n = 10 would be 10 (with exact significance level at 0.04199) or we could use 11 (with an exact significance level of 0.05273).
 - Remember, because we are dealing with ranks, *T*_{crit} must be a discrete number.



T_{calc} versus T_{crit}

http://comp9.psych.cornell.edu/Darlington/wilcoxon/wilcox51.htm

- So, the T_{crit} (labeled S in the table linked above), for n = 10 would be 10 (with exact significance level at 0.04199) or we could use 11 (with an exact significance level of 0.05273).
 - Remember, because we are dealing with ranks, *T*_{crit} must be a discrete number.
- So, since $T_{calc} = 1 < 10 = T_{crit}$ we reject the null hypothesis and conclude that the post-test scores were significantly lower than the pretest scores.



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Current Example applied to *z* test

• Our current example has T = 1 and n = 10 so;



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Chi-square Wilcoxon Kruskal-Wallis Summary

Rank-Sum Test Matched-Pairs Test

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$$z = \frac{T - \frac{n(n+1)}{4}}{\sqrt{\frac{n(n+1)(2n+1)}{24}}} = \frac{1 - \frac{10(10+1)}{4}}{\sqrt{\frac{10(10+1)(2\times10+1)}{24}}} = -2.701$$



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http://www.mathsisfun.com/data/standard-normal-distribution-table.html



 The Kruskal-Wallis test is a nonparametric replacement for the One-way ANOVA when the assumptions of One-way ANOVA are not met.



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 - Both are based on the sums of ranks.
- As with the Wilcoxon's Rank-Sum test we again rank all of the scores (regardless of group membership) and then sum the ranks for each group.

RSS Research and Statistical Support

Omnibus test of differences

• The Kruskal-Wallis test is used to identify differences in central tendency among more than 2 groups.



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Omnibus test of differences

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 - Secondary analysis, such as the Wilcoxon's Rank-Sum test would be necessary (much like conducting post-hoc testing in the ANOVA situation).



• To calculate the Kruskal-Wallis test; compute H



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$$H = \left\lfloor \frac{12}{n_t(n_t+1)} \right\rfloor * \sum \frac{n_t}{n_t} - 3(n_t+1)$$

• where n_t is the total number of participants, R_i is the sum of the ranks in group i, and n_i is the number of participants in group i.



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$$H = \left[\frac{12}{n_t(n_t+1)}\right] * \sum \frac{R_i^2}{n_i} - 3(n_t+1)$$

- where *n_t* is the total number of participants, *R_i* is the sum of the ranks in group i, and *n_i* is the number of participants in group i.
- The comparison distribution is the chi-square distribution with df = k 1 where k is the number of groups.



Quick Example

• Suppose we added limousine drivers to our earlier example comparing driving anger among police officers and taxi drivers.



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Police		Taxi		Limousine	
Score	Rank	Score	Rank	Score	Rank
8	1	27	13	16	9
15	7.5	28	14.5	15	7.5
12	3	19	11	14	6
10	2	17	10	13	4.5
13	4.5	26	12		
		28	14.5		

Tied scores get tied ranks half-way between the two whole number ranks they would occupy if sequential.



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• Then we can calculate H

$$H = \left[\frac{12}{n_t(n_t+1)}\right] * \sum \frac{R_i^2}{n_i} - 3(n_t+1) = \left[\frac{12}{15(15+1)}\right] * \left[\frac{18^2}{5} + \frac{75^2}{6} + \frac{27^2}{4}\right] - 3(15+1) = 11.2275$$

RSS Research and Statistical Support

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Chi-square Wilcoxon Kruskal-Wallis Summary

Summary of Module 11

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Chi-square Wilcoxon Kruskal-Wallis Summary

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Chi-square tests



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Summary of Module 11

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- Chi-square tests
- Wilcoxon's Rank-Sum test
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- Kruskal-Wallis One-Way ANOVA



This concludes Module 11

• Until next time; have a nice day.

These slides initially created on: October 28, 2010 These slides last updated on: November 2, 2010

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